# STUDY OF GOVERNMENT BOND AUCTION IN INDONESIA & THAILAND AS EMERGING COUNTRIES DURING THE PANDEMIC COVID19: AUCTION CYCLE, UNDERPRICING AND COMPETITIVENESS

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#### Abstract

To encourage economic growth, state budgets are typically bigger than anticipated revenues. Due to that reason, to finance these shortfalls (deficits), the government can issue loans bilaterally, multilaterally, or through the issuance of state bonds. When the COVID-19 epidemic expanded broadly and affected the entire globe, the financial markets also exhibited a significant response. Government bonds are the government's principal source of financing, and yield volatility is correlated with debt expenses, therefore the government bond market is large in difficult times. This study examines the empirical study of the auction cycle of government bonds in ASEAN as emerging countries, specifically Indonesia & Thailand and its relationship to the bid-to-cover ratio, exchange rate expectations, local currency, and credit risk pre and during pandemic covid19. According to past research, yields rise toward the auction date, and fall afterward in the secondary market for benchmark tenor samples in each country. The financial data as variables in the market will be using the year 2019-2022 will be processed using the regression data analysis. The study concluded that auction competitiveness reduced secondary market yield on auction day. Yield fluctuation in the secondary market affected by local currency exchange rate forecasts and high credit risk. The modified method from previous research shows a relationship between price changes in the secondary market and when-issued underpricing, indicating that the difference between the auction yield and the secondary market yield is growing, encouraging dealers to quickly secure their profits in the secondary market.

**Keyword:** Auction Cycle, Government Bonds, Underpricing, Competitiveness (Bid to Cover Ratio), Emerging Markets

## Introduction

To achieve economic growth and carry out government operations, a country needs a budget that is typically set annually. This budget consists of revenues and

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expenditures. Revenues come from taxes, duties, grants, and foreign exchange, while expenditures include funds planned by the government for projects such as infrastructure, subsidies, defense costs, government employee salaries and allowances, and so on. Usually, to accelerate economic growth, the government's expenditure budget is larger than the revenue budget. Therefore, to cover budget shortfalls (deficits), the government borrows from other entities through instruments such as bilateral or multilateral loans, or issuing government bonds.

In issuing government bonds, a country can follow policies set by its government. There are several mechanisms commonly used to issue securities, namely through noncompetitive or competitive bidding. In the non-competitive instrument, the government provides opportunities for investors to invest through securities with a private placement mechanism. Meanwhile, through the bidding mechanism, the government usually invites investors, both institutional and retail, from domestic and foreign sources, to compete in providing the best bid price represented by primary dealers and investors can determine their bid prices by referring to the prices of bonds to be auctioned in the secondary market as a benchmark.

According to the efficient market hypothesis, the price of an asset reflects all the information available on that asset, thus eliminating arbitrage opportunities. However, the winning bid price at the auction tends to be lower than the price in the secondary market, which is known as underpricing. The theory of common value auctions has developed over the past 50 years to explain the underpricing phenomenon and help predict the extent of underpricing. Vickrey (1961) initially explained the underpricing phenomenon, and subsequent research by Goldreich (2007) proposed a multi-unit multi-bid auction model as a result of assuming that each auction participant would resell the government bonds they won in the auction, leading participants to anticipate this and engage in bid shading.

Research findings on when-issued underpricing have converged, and in the study by Goldreich (2007), it was found that after the United States switched from a discriminatory to a uniform price auction mechanism, the level of underpricing decreased. This proves that the uniform price mechanism is more efficient in government bond auctions and provides savings in terms of government bond issuance (US Treasury). In addition to research on when-issued underpricing, studies on the secondary market response to auctions have also been a focus of research. Studies by Cammack (1991) and Breedon & Ganley (2000) examined the relationship between underpricing in the whenissued market and bond price movements in the secondary market. The research findings indicate that the secondary market reacts to information received from auctions a few days before and after the auctions take place. Thus, underpricing occurs not only in primary market bond auctions but also in observable patterns of price or yield changes in the secondary market.

Lou et al. (2013) conducted comprehensive research and demonstrated that underpricing is part of a larger recurring cycle triggered by auction participants' anticipation and repeated occurrences of government bond auctions. The research findings showed that the secondary market prices of US Treasury bonds experience a significant decline in the five to ten days leading up to the auction but return to normal within five to ten days after the auction. Lou et al. (2013) used the yield differential of specific tenor US Treasury bonds on day t before or after the auction (Y(t)) compared to the yield in the secondary market at the time of the auction (Y(0)). The graph showing the average data over five days before and five days after the auction forms an inverted V-shaped curve as shown below, indicating that auction participants anticipate the repetitive nature of auctions.

Furthermore, Beetsma et al. (2016a) conducted research in response to price pressure during US Treasury auctions, focusing on Italy and Germany as sample countries in Europe. In their study, Beetsma et al. (2016a) identified price pressure cycles as auction cycles, where bond yields in the secondary market exhibit an inverted V-shaped pattern in the days surrounding the auction, similar to what Lou (2013) found. The research findings indicated that auction cycles occur in Italy with larger and more significant yield movements compared to Germany. Beetsma et al. (2016a) further divided the sample into pre-2007 crisis and post-2007 crisis periods. In German auctions, auction cycles were relatively limited both before and after the crisis, while in Italy, the effects of auction cycles were stronger in the post-crisis period compared to the pre-crisis period.

Therefore, this research aims to examine the empirical study of the auction cycle of government bonds in ASEAN as emerging countries, specifically Indonesia & Thailand and adding its relationship to the bid-to-cover ratio, exchange rate expectations, local currency, and credit risk with yield movement in secondary market during pandemic covid19 era in 2019 - 2022.

## **Research Method**

Specifically, the researchers sampled two countries, Indonesia and Thailand, based on several basic similarities in their economic sectors during the COVID-19 pandemic. All historical data used in the study covers the period from 2019 to 2022, which coincides with the duration of the COVID-19 pandemic and obtained from official Central Bank and Minister of finance website for government bonds auction data and Bloomberg and Refinitiv for daily market data. Both countries have the same method of auction which is the discriminatory for reissuance of government bonds throughout 4 benchmark tenor (5Y, 10Y, 15Y, 20Y). The hypotheses proposed are related to the existence of auction cycles and the relationship between these cycles and the success rate of auctions, exchange rate expectations, and credit risk in the primary market of local currency government bonds.

In this analysis, regression is performed using the ordinary least squares (OLS) method with the Newey-West method to obtain more robust results from the regression of sample averages. We also report confidence intervals at the 90% level to allow for visual interpretation of auction cycles. To test significance, we use t-values within the range of t-5 to t+5 (10 days) as per the research by Lou et al. (2013) and Beetsma et al. (2016a).

## **Results and Discussion Auction Cycle**

The data from the graph depicting the average movements in bond yield changes in the secondary market align consistently with the findings of previous research conducted by Lou et al. (2013) and Beetsma et al. (2016). The graph shows that the average yield changes in the secondary market leading up to the auction day exhibit an increase in yield. However, in Thailand, there is no significant observable downward movement in yield after the government bond auction for the four mentioned maturities. **Bid to cover ratio with secondary market yield movement.** 

Based on the table below, the results of the equation show that the value of  $\alpha$  represents the increase in government bond yield traded in the secondary market during the auction, which is not influenced by the bid-to-cover ratio. The value of  $\beta$  represents the coefficient of the bid-to-cover ratio, where a negative value of  $\beta$  indicates a negative relationship between changes in government bond yield in the secondary market and the bid-to-cover ratio from auction results by investors. The larger the bid-to-cover ratio from the auction results, the smaller the increase in yield in government bonds in the secondary market.

Table 1

			Ir	donesi	a			
	Sample devia	mean tion	Deviation from the previous auction		Deviation from the four- auction average		Deviation from the average bid- to-cover ratio in the previous year	
	Coeffici ent	p- value	Coeffici ent	p- valu e	Coeffici ent	p- value	Coeffici ent	p- value
5 Year α	- 0,00026 7	0.7997 67	- 0,00011 5	0.91 68	0,00001 18	0.991 3	- 0,00013 6	0.9131 58
5 Year β	0,00012 3	0.0006 37 ***	0,00005 6	0.03 07 *	0,00010 1	0.015 5 *	0,00012 8	0.0006 28 ***
Adjust ed R <sup>2</sup>	11,3	6%	4,12%		5,35	%	15,0	8%
10 Year α	- 0.00151 9	0.1636 0	0.00018 99	0.84 7	0.00019 69	0.841	- 0.00021 51	0.8397 93
10 Year β	- 0.00145 8	0.0017 1 **	- 0.00026 31	0.37 4	- 0.00074 40	0.222	- 0.00163 50	0.0001 19 ***
Adjust ed $R^2$	9,52%		0,89%		0,57%		19,03%	
15 Year $\alpha$	0,00105	0.158	0.00101 65	0.16 0	0.00101 83	0.159	0.00145 16	0.118

15 Year β	0,00005 2	0.842	- 0.00018 07	0.35 2	- 0.00030 30	0.332	- 0.00038 41	0.192
Adjust ed $R^2$	0,04	5%	0,989	V0	1,07	%	1,11	%
20 Year α	0,00095 4	0.0702	0,00097	0.06 32 .	0,00096	0.066 1.	0,00065	0.313
20 Year β	- 0,00000 9	0.9224	0,00005 8	0.48 36	- 0,00001 1	0.929 8	-0,00003	0.766
Adjust ed $R^2$	0,01	.%	0,569	V <sub>0</sub>	0,009	)%	0,14	%

Note: The estimation method used is Ordinary Least Square (OLS) with fixed effects. The symbols •, \*, \*\*, \*\*\* indicate significance levels in ascending order at 90%, 95%, 99%, and 99.9% respectively.

			]	Table 2 Fhailand	l			
	Sample deviat	mean ion	Deviation from the previous auction		Deviation from the four-auction average		Deviation from the average bid-to-cover ratio in the previous year	
	Coeffici ent	p- valu e	Coeffici ent	p- value	Coeffici ent	p- value	Coeffici ent	p- valu e
5 Year $\alpha$	0.01071 5	0.07 15 .	0.011093	0.063 8 .	0.01071 5	0.063 5 .	0.01536 3	0.04 35 *
5 Year β	0.01363 0	0.06 68 .	- 0.01208 7	0.027 7 *	- 0.01636 8	0.028 9 *	- 0.011907	0.13 81
Adjust ed R <sup>2</sup>	10.12	%	16.54%		15.58	3%	6.89%	V <sub>0</sub>
10 Year α	0.00521 8	0.12 38	0.00494 4	0.206	0.00517 4	0.186	0.00315 8	0.50 5
10 Year β	0.00949 1	0.03 32 *	0.00460 2	0.232	0.00128 5	0.802	- 0.00244 7	0.64 4
Adjust ed $R^2$	21.94	%	3.61	%	0.4%		2.489	V <sub>0</sub>
15 Year $\alpha$	0.00325 2	0.54 81	0.001135	0.728 07	0.00325 2	0.281 76	0.01468 1	0.40 9
15 Year $\beta$	- 0.01504 9	0.01 62 *	- 0.00947 6	0.004 05 **	- 0.01798 4	0.001 19 **	0.01608	0.22 0

Adjust ed $R^2$	74.94	%	94.06	5%	93.04	1%	77%	)
20 Year α	0.00360 3	0.30 67	0.00506 5	0.215	0.00319 8	0.441	0.00450 3	0.31 9
20 Year β	- 0.00837 9	0.07 75 .	- 0.004011	0.275	- 0.00292 4	0.696	- 0.00796 4	0.12 3
Adjust ed $R^2$	18.86	%	2.94	%	1.44	%	17.95	%

Note: The estimation method used is Ordinary Least Square (OLS) with fixed effects. The symbols •, \*, \*\*, \*\*\* indicate significance levels in ascending order at 90%, 95%, 99%, and 99.9% respectively.

In this study, significant regression results were found for the 5-year tenor in both Indonesia and Thailand. These findings are consistent with Forest's (2012) research, which indicates that regression against deviations in the bid-to-cover ratio in US government bond auctions is significantly related to yield movements in the secondary market for the 5-year tenor. The regression results are also consistent with the findings of Beetsma et al. (2018) for the 5-year and 10-year tenors, showing partial consistency for the 15-year tenor but inconsistency for the 20-year tenor.

			]	Cable 3				
			In	donesia	a			
	Sample mean deviation		Deviation from the previous auction		Deviation from the four- auction average		Deviation from the average bid- to-cover ratio in the previous year	
	Coeffici ent	Coeffici p- ent value		p- valu e	Coeffici ent	p- valu e	Coeffici ent	p- value
5 Year α	- 0,00058 4	0.596 87	-0,00042	0.71 71	- 0,00029 5	0.79 59	-0,00054	0.686 40
5 Year β	0,00012	0.001 02 **	0,00005 2	0.05 26 .	0,00009 6	0.02 57 *	0,00013	0.001 04 **
5 Year γ	0,026	0.499 26	0,02	0.62 21	0,023	0.56 81	0,01654	0.688 72
Adjust ed $R^2$	11,03	3%	3,169	%	4,549	%	14,37	7%
10 Year α	- 0.00215 9	0.078 24 .	0.00052 36	0.63 6	- 0.00048 39	0.66 1	- 0.00055 52	0.666 55

Bid to cover ratio with secondary market yield movement & 1Y forward rate.

10 Year β	- 0.00148 5	0.005 03 **	- 0.00034 98	0.29 5	- 0.00069 68	0.34 2	- 0.00160 07	0.001 99 **
10 Year γ	0.00146 84	0.978 42	- 0.02439 75	0.66 9	- 0.02437 70	0.67 0	0.00151 97	0.980 02
Adjust ed R <sup>2</sup>	7,72	%	1,569	%	0,979	%	14,27	7%
15 Year α	0.00057 2	0.514	0.00041 37	0.60 5	- 0.00043 98	0.58 2	- 0.00058 11	0.616
15 Year β	- 0.00015 4	0.664	- 0.00030 23	0.28 9	- 0.00051 53	0.24 7	- 0.00034 66	0.418
15 Year γ	0.06053 18	0.266	0.05902 21	0.27 5	0.05908 34	0.27 3	0.07161 00	0.300
Adjust ed R <sup>2</sup>	2,09	%	0,589	%	0,899	%	3,939	%
20 Year α	0.00035 2	0.647 7	0,00054	0.34 9	0.00052 53	0.36 5	0.00029 53	0.701 6
20 Year β	- 0.00059 1	0.090 6 .	- 0,00008 1	0.78 7	- 0.00028 78	0.52 4	- 0.00066 70	0.066 2.
20 Year γ	0.03970 58	0.305 6	0,0414	0.30 4	0.03775 58	0.35 1	0.031140 5	0.514 8
Adjust ed $R^2$	3,33	%	2,039	%	2,559	%	5,089	%

Note: The estimation method used is Ordinary Least Square (OLS) with fixed effects. The symbols •, \*, \*\*, \*\*\* indicate significance levels in ascending order at 90%, 95%, 99%, and 99.9% respectively.

			r	<b>Fable</b> 4				
			Τ	'hailano	d			
	Sample deviat	mean ion	Deviation from the previous auction		Deviation the four-a avera	n from nuction age	Deviation from the average bid- to-cover ratio in the previous year	
	Coefficie nt	p- valu e	Coefficie nt	p- valu e	Coefficie nt	p- value	Coefficie nt	p- valu e
5 Year $\alpha$	0.008783	0.118 0	0.009015	0.10 49	0.008935	0.105 6	0.012377	0.09 01 .
5 Year	-	0.12	-	0.03	-	0.073	-	0.19
β	0.010781	64	0.010883	25 *	0.012952	9.	0.009808	92

5 Year γ	0.043040	0.05 34 .	0.044483	0.03 77 *	0.039636	0.071 8 .	0.042526	0.09 25 .
Adjust ed $R^2$	21%	)	29.15	%	24.09	%	16.959	%
10 Year $\alpha$	0.005332	0.13 15	0.005450	0.18 1	0.005276	0.197	0.003853	0.45 1
10 Year β	0.009489	0.03 95 *	0.005732	0.18 0	0.001072	0.843	- 0.002444	0.65 7
10 Year γ	0.006299	0.78 35	0.019819	0.47 6	0.005265	0.847	0.015978	0.56 9
Adjust ed R <sup>2</sup>	16.83	%	0.33%	6	0.719	%	6.59%	6
15 Year α	0.002678	0.59 78	0.000983 8	0.83 01	0.003482	0.289 97	0.000810 6	0.87 94
15 Year β	- 0.015484	0.02 07 *	- 0.009452 2	0.02 34 *	- 0.018592	0.004 28 **	- 0.016719 9	0.08 82 .
15 Year γ	- 0.020346	0.27 27	0.001326	0.93 88	0.008158	0.451 20	0.052905 1	0.17 42
Adjust ed R <sup>2</sup>	79.1%	⁄0	91.129	%	92.57	°%	94.479	%
20 Year $\alpha$	0.002398	$\begin{array}{c} 0.50 \\ 0 \end{array}$	0.003745	0.37 0	0.001588	0.708	0.003621	0.46 5
20 Year β	- 0.008171	0.08 1 .	- 0.004105	0.26 2	- 0.004433	0.554	- 0.007169	0.19 6
20 Year γ	- 0.012675	0.24 7	- 0.012740	0.29 4	- 0.014734	0.249	- 0.007418	0.58 2
Adjust $P^2$	22.46	%	5.24%	6	14.27	%	10.499	%

Note: The estimation method used is Ordinary Least Square (OLS) with fixed effects. The symbols •, \*, \*\*, \*\*\* indicate significance levels in ascending order at 90%, 95%, 99%, and 99.9% respectively.

In this context,  $\alpha$  represents the direct impact of the auction itself on yields in the secondary market.  $\beta$  depicts the influence of auction competitiveness on yield changes in the secondary market, with higher levels of competition exerting greater pressure on yields. Meanwhile,  $\gamma$  reflects the impact of fluctuations in the USD-to-local currency exchange rate on yields in the secondary market. The regression results show that by including the variable of changes in the forward exchange rate of the local currency to USD with a 1-year tenor, this model reveals an increase in the t-statistics for the  $\beta$  coefficient. The magnitude of the coefficient remains relatively consistent across the regression results for the 5-year and 10-year tenors in both sampled countries.

In both sampled countries, the  $\alpha$  level, which is relatively similar to the previous regression, indicates consistency in positive yield changes in the secondary market on the day of the auction. The  $\beta$  coefficient also remains consistent with a negative and consistently significant value, indicating that higher levels of competition result in smaller yield changes in the secondary market. Furthermore, the  $\gamma$  value is positive, indicating that an increase in the USD value relative to the local currency or a weakening of the local currency against the USD leads to an increase in forward rates. This, in turn, impacts an increase in yields in the secondary market on the day of the auction.

			т	Table 5				
	Sample devia	e mean tion	mean ion ion ion ion ion ion ion ion ion io		a Deviation the fo auction a	n from ur- verage	Deviatio the avera to-cover the pre	on from age bid- ratio in evious ar
	Coeffici ent	p- value	Coeffici ent	p- value	Coeffici ent	p- value	Coeffici ent	p- value
5 Year $\alpha$	0,00032	0.7937	0,00018	0.891 4	0.00046 81	0.715 31	0.00027 24	0.8510
5 Year β	0,00016	0,0000 87 ***	0,00006 4	0.022 3 *	0.00012 55	0.006 73 **	0.00015 92	0.0002 42 ***
5 Year θ	0,0712	0.0305 *	0,035	0.289 5	$\begin{array}{c} 0.04787\\70\end{array}$	0.153 06	0.06179 65	0.0824 45 .
Adjust ed $R^2$	17,1	2%	4,64	%	7,38	%	19,3	8%
10 Year α	- 0.00186 3	0.1502	- 0.00040 94	0.731 0	- 0.00039 94	0.738 3	- 0.00051 86	0.7085
10 Year β	- 0.00131 6	0.0109 *	- 0.00037 54	0.242 3	- 0.00068 17	0.336 4	- 0.00133 84	0.0089 **
10 Year <i>θ</i>	0.06654 72	0.0259 *	0.07407 64	0.017 2 *	0.07147 64	0.021 9 *	0.05675 52	0.0732
Adjust ed $R^2$	14,4	6%	7,28	%	6,63	%	16,9	9%
15 Year α	0.00062 51	0.5207 1	0,00021	0.807 27	0,00019 6	0.819 05	- 0.00035 12	0.7691
15 Year β	0.00037 24	0.3800 1	- 0,00000 09	0.997 56	- 0,00008 8	0.867 07	0.00016 34	0.7593
15 Year θ	0.06567 31	0.0045 2 **	0,067	0.003 69 **	0,068	0.003 63 **	0.06292 41	0.0163 *

Bid to cover ratio with secondary market yield movement & credit default swap.

Adjust ed $R^2$	12,05	5%	10,84	1%	10,88	8%	10,09	9%
20 Year $\alpha$	- 0.00041 2	0.650	0.00068 89	0.306	0.00060 7	0.366	0.00032 71	0.713
20	-		-		-		-	
Year $\beta$	0.00063	0.110	0.00040	0.221	0.00047	0.316	0.00065	0.110
	9		58		54		40	
20	0.01736	0.200	0.01995	0.256	0.01813	0.200	0.01029	0 5 9 2
Year $\theta$	37	0.309	31	0.230	60	0.299	78	0.385
Adjust ed R <sup>2</sup>	2,69	9%	0,72	%	3,43	%	2,80	%

Note: The estimation method used is Ordinary Least Square (OLS) with fixed effects. The symbols •, \*, \*\*, \*\*\* indicate significance levels in ascending order at 90%, 95%, 99%, and 99.9% respectively.

			Τ	able 6				
			T	hailand				
	Sample deviat	mean ion	Deviation from the previous auction		Deviation from the four-auction average		Deviation from the average bid-to-cover ratio in the previous year	
	Coeffici ent	p- value	Coeffici ent	p- valu e	Coeffici ent	p- value	Coeffici ent	p- valu e
5 Year $\alpha$	0.011468	0.150 6	0.014816	0.07 41 .	0.009560	0.216 9	0.01790	0.10 1
5 Year β	- 0.021966	0.052 4 .	- 0.012639	0.05 02 .	- 0.026529	0.024 4 *	-0.02144	0.13 2
5 Year θ	0.180161	0.347 5	0.186248	0.34 07	0.153834	$\begin{array}{c} 0.404 \\ 0 \end{array}$	0.22544	0.34 3
Adjust ed R <sup>2</sup>	17.08	s%	18.27	%	23.74	%	12.97	%
10 Year α	0.005489	0.076 40 .	0.004493	0.30 4	0.004900	0.276	- 0.001817	0.84 7
10 Year β	0.014240	0.001 98 **	0.006304	0.118	0.00551	0.424	- 0.004738	0.71 7
10 Year <i>θ</i>	0.268652	0.034 57 *	0.159423	0.311	0.198867	0.316	0.139998	0.52 8
Adjust ed $R^2$	54.28	3%	13.13	%	9,849	%	18,479	%
15 Year $\alpha$	0.004589	0.527 6	0.000912 7	0.83 9	0.003858	0.334 14	0.003000	0.81 3
15 Year β	0.016236	0.050 2 .	- 0.009387 4	0.02 7 *	- 0.018484	0.007 05 **	0.012052	0.23 1

15 Year θ	0.128418	0.706 5	- 0.023129 5	0.90 2	0.058250	0.730 88	0.108330	0.83 5
Adjust ed $R^2$	68.4	%	91.17	%	91.14	%	32.389	%
20 Year α	0.003989	0.510	0.006899	0.37 1	0.006166	0.295	- 0.005346	0.73 5
20 Year β	- 0.009050	0.334	- 0.002770	$\begin{array}{c} 0.78\\ 8\end{array}$	- 0.012116	0.309	- 0.029783	0.24 9
20 Year θ	0.046134	0.915	- 0.118416	0.78 7	- 0.080556	0.838	- 0.097345	0.93 9
Adjust ed $R^2$	16.84	%	3.43%	⁄0	18.35	5%	0.72%	6

Note: The estimation method used is Ordinary Least Square (OLS) with fixed effects. The symbols •, \*, \*\*, \*\*\* indicate significance levels in ascending order at 90%, 95%, 99%, and 99.9% respectively.

The positive coefficient  $\theta$  in both sampled countries indicates that when the credit risk of the issuing country increases, the yield changes in the secondary market on the day of the auction also become larger. The regression analysis findings suggest that when the variable of credit risk changes, measured by the level of Credit Default Swap (CDS) 5-Year, is included, a significant portion of the  $\beta$  coefficients in the auction model with a 5-Year, 10-Year, and 15-Year tenor exhibit significant levels of significance different from zero. This indicates a significant relationship between changes in credit risk and yield changes in the secondary market in the context of auctions for bonds with these tenors. Additionally, it suggests that yields in the secondary market reflect changes in the credit risk of the issuing country.

Overall, the significance of the coefficient for the credit risk variable is consistent with the findings of Beetsma et al. (2018) and Gonzalez-Hermosillo (2008), which state that CDS serves as a proxy for representing the credit risk of a country. Additionally, the significance of the coefficient for the competition variable confirms the findings of previous regressions, which suggest that competition level has a significant influence on the yield changes of government bonds in the secondary market.

Bid to cover ratio with secondar	y market yield movement	& U	ınderpricin	ıg
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	Table 7	1
	Secondary M	larket
Indonesia	UP=	= yA - yT
Indonesia	Coefficient	p-value
5 Year $\lambda$	0.003135	0.00697 **
5 Year μ	-0.133186	0,0000006 ***
Adjusted R <sup>2</sup>	2	3,54%
10 Year $\lambda$	0.0031862	0.00138 **

10 Year $\mu$	-0.1487679	0,00000004 ***
Adjusted $R^2$		28,01%
15 Year $\lambda$	0.00195	0.01330 *
15 Year μ	-0.044883	0.00814 **
Adjusted R <sup>2</sup>		6,64%
20 Year $\lambda$	0.0011703	0.0254 *
20 Year $\mu$	-0.0089590	0.0475 *
Adjusted R <sup>2</sup>		3,34%
Thailand	UF	P = yA - yT
Thananu	Coefficient	p-value
5 Year $\lambda$	0.010203	0.0803 .
5 Year μ	-0.071053	0.4773
Adjusted R <sup>2</sup>		2.22%
10 Year $\lambda$	-0.001760	0.859
10 Year $\mu$	0.462611	0.049 *
Adjusted R <sup>2</sup>		18.33%
15 Year $\lambda$	0.003544	0.601
15 Year μ	0.001607	0.989
Adjusted R <sup>2</sup>	(	0.0005%
20 Year $\lambda$	-0.0006853	0.887
20 Year $\mu$	0.0433277	0.416
Adjusted $R^2$		6,10%

Based on the above regression results, it can be observed that the yield changes of government bonds in the secondary market during the auction day have a significant relationship with underpricing measurements in Indonesia for all tenors at confidence levels ranging from 95% to 99%. In Thailand, this relationship is significant only for the 5-year tenor at a 90% confidence level, aligning with the research methodology of Goldstein (1962) and Cammack (1991).

In regressions with significant  $\mu$  values, there is also a positive and significant  $\lambda$  value for Indonesia across all four tenors at confidence levels ranging from 95% to 99%, while for Thailand, it is significant only for the 10-year tenor at a 95% confidence level. This indicates that the yield changes of government bonds in the secondary market during the auction day have a larger magnitude compared to the yield differences obtained in the auction relative to the previous day's closing yield or the average yield of government bonds in the secondary market on the day before and during the auction.

Furthermore, the influence of competition level on the changes in bond yields in the secondary market within the underpricing control framework in the equation above is demonstrated through regression analysis with the following results:

				Table 8				
	Sampl devi	e mean ation	Deviatio the pro auct	Indonesia on from evious tion	a Deviati the four ave	on from -auction rage	Deviatio the avo bid-to- ratio i previou	n from erage cover n the is year
	Coeffic ient	p-value	Coeffic ient	p- value	Coeffic ient	p-value	Coeffic ient	p- value
5 Year α	0,0025	0.0342 *	0,0028	0.0155 *	0,0029	0.0128 *	0,0027	0.054 67 .
5 Year ß	0,0000 58	0.1051	0,0000 28	0.2327	0,0000 43	0.2699	0,0000 63	0.097 314 .
γ 5 Year κ	-0,113	0,00008 ***	-0,125	0,0000 049 ***	-0,124	0,00000 821 ***	-0,11	0.000 547 ***
Adjus ted R <sup>2</sup>	24,9	95%	23,7	9%	23,	74%	28,3	5%
10 Year α	0.0017 566	0.1191	0.0031 394	0.0018 1 **	0.0031 131	0.00176 **	0.0026 384	0.016 06 *
10 Year β	- 0.0009 57	0.0196 *	- 0.0001 02	0.6848 3	- 0.0005 99	0.24714	0.0011 035	0.002 94 **
10 Year κ	- 0.1352 86	0,00000 045 ***	- 0.1476 68	0,0000 007 ***	- 0.1474 00	0,00000 005 ***	0.1337 72	0,000 004 ***
Adjus ted <i>R</i> <sup>2</sup>	31,:	58%	27,3	3%	28,2	29%	40,9	3%
15 Year α	0.0020 520	0.01224 *	0.0019 348	0.0145 5 *	0.0019 331	0.01463 *	0.0019 853	0.039 3 *
15 Year β	0.0001 220	0.62651	- 0.0001 49	0.4278 3	- 0.0002 44	0.42015	- 0.0002 69	0.360 6
15 Year к	- 0.0457 26	0.00762 **	0.0440 38	0.0096 7 **	- 0.0438 83	0.00996 **	0.0377 82	0.067 4 .

Adjus								
ted	5,8	33%	6,23	5%	6,2	8%	4,71	%
$R^2$								
20	0.0020	0.01224		0.0252		0.0261	0 0008	
Year	520	0.0122 <del>4</del> *	0,0012	0.02 <i>32</i> *	0,0012	*	6	0.190
α	520						0	
20	0.0001		0 0000		0 0000		-	
Year	220	0.62651	0,0000 57	0.4912	15	0.9015	0,0000	0.947
β	220		57		15		07	
20	-	0.00762		0.0401	-	0.0496		
Year	0.0457	0.00702 **	-0,0089	0.0491 *	0,0090	0.0 <del>4</del> 60 *	-0,0074	0.136
κ	26				2			
Adjus								
ted	5,8	33%	2,70	6%	2,2	4%	0,57	'%
$R^2$								

Note: The estimation method used is Ordinary Least Square (OLS) with fixed effects. The symbols •, \*, \*\*, \*\*\* indicate significance levels in ascending order at 90%, 95%, 99%, and 99.9% respectively.

			Т	Table 9				
			T	hailand				
	Sample mean deviation		Deviation the prev auction	n from vious on	Deviation the four-a avera	n from Juction ge	Deviation the aver bid-to-co ratio in previous	from age over the year
	Coefficie nt	p- value	Coefficie nt	p- value	Coefficie nt	p- value	Coefficie nt	p- valu e
5 Year α	0.011573	0.084 1 .	0.011296	0.064 6 .	0.012070	0.068	0.002187	0.71 1
5 Year β	- 0.002777	0.647 2	- 0.004205	0.282 4	- 0.004024	0.508	- 0.002927	0.99 5
5 Year κ	- 0.104284	0.405 8	- 0.124492	$\begin{array}{c} 0.255\\ 0\end{array}$	- 0.116329	0.343	- 0.001390	0.89 1
Adjust ed R <sup>2</sup>	3.129	%	7.9%	6	4.19%		0.16%	
10 Year α	- 0.001794	0.860 6	- 0.000204 7	0.985 2	- 0.000756 3	0.940 3	0.003405	0.82 3
10 Year β	0.002691	0.788 5	- 0.003495 1	0.674 6	0.008431 8	0.440 3	0.015183	0.28 5
10 Year κ	0.463794	0.056 3.	0.442151 2	0.079 9 .	0.417106 7	0.086 2.	0.487185	0.13 8
Adjust ed $R^2$	12.96	%	9.979	%	16.27	%	16.949	%

15 V	0.003263	0.705	0.005087	0.641	0.003661	0.676	0.004351	0.64
Year $\alpha$	6				/			3
15	-	0.033	-	0 976	0.000244	0 97/	-	0.69
Year $\beta$	0.000557	0.755	0.000136	0.770	7	0.774	0.002437	3
15	0 011 475				-			0.75
Year <i>k</i>	0.0114/5	0.950	0.024790	0.912	0.002514	0.989	0.058307	0.75
-	6				6			2
Adjust		. ,	4		0.0.	. ,	10.010	. ,
, _ <b>,</b> _ <b>,</b>	0.289	/0	1.029	0	0.059	20	10.319	/
ed $R^2$	0.207	•	1.02	0	0.00	0	10.017	
ed $R^2$ 20	0.003263	0.705	-	0.((2	-		-	0.61
$\frac{\text{ed } R^2}{20}$ Year $\alpha$	0.003263	0.705	0.002251	0.663	- 0.001079	0.836	- 0.003646	0.61
ed $R^2$ 20 Year $\alpha$ 20	0.003263	0.705	0.002251	0.663	0.001079	0.836	0.003646	0.61 6 0.56
ed $R^2$ 20 Year $\alpha$ 20 Year $\beta$	0.003263 6 - 0.000557	0.705 0.933	0.002251 0.004524	0.663 0.227	0.001079	0.836 0.752	0.003646	0.61 6 0.56 5
ed $R^2$ 20 Year $\alpha$ 20 Year $\beta$ 20	0.003263 6 0.000557 0.011475	0.705 0.933	0.002251	0.663 0.227	0.001079	0.836 0.752	0.003646	0.61 6 0.56 5 0.28
ed $R^2$ 20 Year $\alpha$ 20 Year $\beta$ 20 Year $\kappa$	0.003263 6 - 0.000557 0.011475 6	0.705 0.933 0.950	0.002251 0.004524 0.063658	0.663 0.227 0.274	0.001079 0.002233 0.044677	0.836 0.752 0.424	0.003646 0.003234 0.125248	0.61 6 0.56 5 0.28 8
ed $R^2$ 20 Year $\alpha$ 20 Year $\beta$ 20 Year $\kappa$ Adjust	0.003263 6 0.000557 0.011475 6	0.705 0.933 0.950	0.002251 0.004524 0.063658	0.663 0.227 0.274	0.001079 0.002233 0.044677	0.836 0.752 0.424	0.003646 0.003234 0.125248	0.61 6 0.56 5 0.28 8

Note: The estimation method used is Ordinary Least Square (OLS) with fixed effects. The symbols •, \*, \*\*, \*\*\* indicate significance levels in ascending order at 90%, 95%, 99%, and 99.9% respectively.

Based on the previously mentioned regression results, it can be observed that the competition level still has a significant influence on the changes in bond yields in the secondary market on auction days. In this context, the variable of underpricing is used, measured based on the method proposed by Cammack (1991). The analysis results indicate that underpricing has a significant effect on the changes in bond yields in the secondary market on auction days for all four tenors in Indonesia, but not significant in Thailand.

However, there are some differences in both countries with negative coefficient values  $\kappa$ , indicating that underpricing acts as a reducing factor for the changes in bond yields in the secondary market on auction days. In other words, the larger the underpricing, the greater the changes in bond yields in the secondary market. When connecting it with the previously mentioned regression results, the negative value of  $\kappa$  provides an interpretation that a larger underpricing, which can be explained by increased information dispersion, actually reduces the magnitude of changes in bond yields in the secondary market. However, the  $\kappa$  value shows significance in Indonesia, indicating that the difference in yields achieved in auctions compared to the closing yield the day before affects the changes in bond yields in the secondary market on auction days.

## **Conclusion:**

The research conducted on the pattern of yield movements in the secondary market before and after auctions (auction cycle) indicates that in Indonesia and Thailand, there is an increasing trend in yields five days before the auction. Although there is variation in the pattern of yield decreases five days after the auction, most research findings indicate a decline in yields after the auction takes place.

When the demand for the won bonds increases, the competition to acquire those bonds also intensifies. As a result, auction participants are motivated to make purchases from the secondary market to meet the demand. However, this relationship cannot be consistently explained in benchmark bond auctions with a 20-year tenor, as well as in some other regressions involving those benchmark bonds. This is due to irregular auction frequency and long intervals between auctions.

Based on the proposed hypothesis, there is a relationship between expectations of the local currency exchange rate and competition level in auctions with changes in government bond yields in the secondary market on the auction day. If there is an expectation that the local currency will depreciate against the US Dollar (USD), the increase in yields in the secondary market will be greater. Conversely, if there is an expectation of currency appreciation, the increase in yields in the secondary market will be reduced.

Based on the proposed hypothesis, there is a relationship between credit risk level and competition level in auctions with changes in government bond yields in the secondary market on the auction day. Increasing credit risk level becomes a consideration for investors in the secondary market to increase the risk premium on market yields. In this context, the higher the credit risk level associated with government bonds, the higher the expected risk premium by investors. This risk premium reflects the additional compensation required by investors in return for the possibility of default or the government's inability to meet bond payment obligations. Thus, credit risk level can influence changes in government bond yields in the secondary market on the auction day. Investors will tend to demand higher yields to compensate for higher credit risk, leading to an increase in yields in the secondary market.

The research findings regarding the relationship between changes in government bond yields in the secondary market on the auction day and when-issued underpricing, using the measurement method by Cammack (1991), yield results that contradict common intuition. In auctions with a 10-year tenor, there is a significant negative impact between changes in yields in the secondary market and the level of underpricing, which can be explained by information dispersion according to Cammack (1991).

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